Part I

Probabilistic Method

WHICH IS THE MOST BEAUTIFUL EQUATION?

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$e^{i\pi} + 1 = 0$

How to prove that some object exists?

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SOME OBJECT EXISTS?

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- To prove that probability that such an object exists is positive a non constructive approach.

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Their power comes from our ability to reformulate, in various ways, so called **counting arguments** in the language of probability and then *to apply* various tools of the probability theory.

Example: One can show that for every $n \times n$ 0-1-matrix A, and for any randomly chosen vector $b \in \{-1, +1\}^n$, it holds

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From that we may conclude that for every such a matrix A, there always exists a vector $b \in \{-1, +1\}^n$ such that

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- If such a probability is indeed large then we can find such an object quite efficiently just by applying a random searching process - a sampling experiment.
- In some cases, however, no explicit construction of a combinatorial object is known in spite of the fact that we can show that such object exists.

EXAMPLE

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- The random process P is then analyzed and some conclusions are made that are, or at least look as, independent of the experiment \mathcal{E} .

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That implies that there must be a partition satisfying the theorem.

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- In some other cases the existence proof obtained by the probabilistic method can be converted even to an efficient deterministic algorithm to find a desirable object *O*such a process is called derandomization.

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Since we can test in polynomial time whether the value of the cut determined by a sample is at least m/2, by counting edges crossing the cut, we have a Las Vegas algorithm to find a cut.

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we get p = 1/(k + 1). For this p and large k, the expected number of hats is

$$np(1-p)^k = n\left(\frac{1}{k+1}\right)\left(1-\frac{1}{k+1}\right)^k \leq \frac{n}{(k+1)e^k}$$

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By connecting these two paths through the node v we get a Hamiltonian path for the tournament T.

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Theorem now follows from the following calculations:

$$\mathbf{E}[X] = \mathbf{E}[\sum_{\sigma} X_{\sigma}] = \sum_{\sigma} \mathbf{E}[X_{\sigma}] = n! (\frac{1}{2})^{n-1}.$$

IV054 1. Probabilistic Method

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Therefore, with a positive probability no event A_K occurs. That is, there is a tournament on *n* vertices that has the property S_k .

EXPLANATION

If K is a set on k players, then the probability that a player P not in K beats all of them is 2^{-k} If K is a set on k players, then the probability that a player P not in K beats all of them is 2^{-k} and the probability is $1 - 2^{-k}$ that he does not beat all of them. If K is a set on k players, then the probability that a player P not in K beats all of them is 2^{-k} and the probability is $1 - 2^{-k}$ that he does not beat all of them.

Since there are n - k players outside the group K, the probability that none of them beats all players in K is $(1 - 2^{-k})^{n-k}$.

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Therefore, there exists at least one assignment for which $\sum_{i=1}^{m} Z_i \geq \frac{m}{2}$.

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If A is a randomized algorithm, then $m_A(I)$ is a random variable and in such a case $m_A(I)$ is replaced by $\mathbf{E}[m_A(I)]$ in the definition of the performance ratio.

We now show the existence of a randomized algorithm for MAXSAT with performance ratio $\frac{3}{4}$.

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We now show another algorithm that performs especially well when there are (many) clauses consisting of a single literal.

Finally, we show that on any input instance, one of the two designed algorithms yields a randomized $\frac{3}{4}$ -approximation algorithm.

BASIC IDEA

Reformulate the problem as a 0-1 linear programming problem.

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Notation: With each clause C_j , in the given input formula, we associate an indicator variable $c_j \in \{0, 1\}$, that indicates whether or not the clause C_j is satisfied at the algorithm being used.

INDICATOR VARIABLES

Moreover, to each variable x_i we assign an indicator variable v_i defined by

$$x_i = true \iff v_i = 1$$

 C_j^+ - set of indices of variables that appear uncomplemented in C_j C_i^- - set of indices of variables that appear complemented in C_j

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$$v_i, c_j \in \{0,1\} \quad (\forall i,j) \qquad (\star)$$

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Let \hat{v}_i (\hat{c}_i) be the value of variable v_i (c_i) obtained by solving the rational linear programming problem. Clearly, $\sum_{i=1}^{n} c_i \leq \sum_{j=1}^{n} \hat{c}_j$.

We first show that using the randomized rounding method we obtain a truth assignment for which the expected number of satisfied clauses is at least $(1 - \frac{1}{e}) \sum_{j} \hat{c}_{j}$.

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This will follow from the Lemma shown on the next slide for the case we use the following **randomized rounding**: each v_i is set, independently, to 1 with the probability \hat{v}_i .

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Notation: For an integer k denote $\beta_k = 1 - (1 - \frac{1}{k})^k > 1 - \frac{1}{e}$.

Lemma: Let C_j be a clause with k literals. The probability that it is satisfied by the randomized rounding is at least $\beta_k \hat{c}_j > (1 - \frac{1}{e})\hat{c}_j$.

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Since function $f(x) = 1 - (1 - \frac{x}{k})^k$ is concave, it suffices to verify the above inequality for x = 0 and x = 1 what is easy.

From the last Lemma, and from the linearity of expectations, it follows:

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Theorem: Given an instance of MAX-SAT, the expected number of clauses satisfied by linear programming and randomized rounding is at least $(1 - \frac{1}{e})$ time the maximum number of clauses that can be satisfied on that instance.

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k	$1 - 2^{-k}$	β_k
1	0.5	1.0
2	0.75	0.75
3	0.875	0.704
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We now show that on any instance one of the algorithms is a $\frac{3}{4}$ - approximation algorithm for the MAX-SAT problem:

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Let n_2 denote the expected number of clauses that are satisfied when we use the linear programming followed by the randomized rounding (what corresponds to the second algorithm).

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Let S_k denote the set of clauses that contain k literals. We know that

$$egin{aligned} n_1 &= \sum_k \sum_{C_j \in \mathcal{S}_k} (1-2^{-k}) c_j \geq \sum_k \sum_{C_j \in \mathcal{S}_k} (1-2^{-k}) \hat{c}_j, \ n_2 &\geq \sum_k \sum_{C_j \in \mathcal{S}_k} eta_k \hat{c}_j. \end{aligned}$$

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Since $(1-2^{-k}) + \beta_k \geq \frac{3}{2}$ for all k, we get

$$\frac{n_1 + n_2}{2} \geq \frac{3}{4} \sum_k \sum_{C_j \in S_k} \hat{c}_j = \frac{3}{4} \sum_j \hat{c}_j.$$

RAMSEY NUMBER PROBLEM

The **Ramsey number** R(k, l) is the smallest integer n such that in any 2-coloring of the edges of a complete graph K_n , on n nodes, by red and blue, there either is a red K_k (i.e. a complete subgraph on k vertices with edges coloured red), or there is a blue K_l .

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Ramsey (1930) showed that R(k, l) is finite for any two integers k and l.

We use the probabilistic method to show a lower bound on R(k, k).

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$$\binom{n}{k} 2^{1-\binom{k}{2}} < \frac{2^{1+k/2}}{k!} \frac{n^k}{2^{k^2/2}} < 1$$

and hence $R(k, k) > 2^{k/2}$ for all $k \ge 3$.

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Since there are $2^{\binom{n}{2}}$ possible colorings, an exhaustive search cannot be efficient.

However, a closer look at the proof of the last theorem shows that the proof can be used to produce effectively a coloring that is very likely to be good. This is due to the fact for large k if $n = \lfloor 2^{k/2} \rfloor$, then

$$\binom{n}{k} 2^{1-\binom{k}{2}} < \frac{2^{1+\frac{k}{2}}}{k!} (\frac{n}{2^{k/2}})^k \le \frac{2^{1+\frac{k}{2}}}{k!} << 1$$

because $\binom{n}{k} \leq \frac{n^k}{k!}$. Hence, a random coloring of K_n is very likely not to contain a monochromatic $K_{2 \lg_2 n}$.

As a consequence of previous results, if we need to find a two-coloring of edges of K_{1024} without a monochromatic K_{20} we can simply produce a random two-coloring and then the probability that it contains a monochromatic K_{20} is less than $\frac{2^{11}}{201}$ what is much, much less than probability of error in any proof that a certain coloring is good.

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In 1993 S. P. Radziszowski and B. D. McKay showed that R(4,5) = 25. They estimate that their computer proof consumed an equivalent of 11 years of computation by a standard desktop computer.

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- It is first shown that with a positive probability an object O' exists that is very close, in some sense, to O.
- Secondly, O' is changed, to obtain O, and it is shown that the probability of the existence of O remains positive.

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Let (p,q) be the base of the triangle (p,q,r) and let ||p-q|| = b.

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$$\begin{aligned} \Pr[A(p,q,r) \leq \varepsilon] &= \int_{b=0}^{\sqrt{2}} \Pr[b \leq ||p-q|| \leq b + \Delta b] \times \Pr[\text{triangle. } h. \leq \frac{2\varepsilon}{b}] \\ &\leq \int_{b=0}^{\sqrt{2}} \frac{2\sqrt{2}\varepsilon}{b} 4\pi b\Delta b = 16\pi\varepsilon. \end{aligned}$$

Let us now compute the expected number of triangles with the area $\leq \varepsilon = \frac{1}{100n^2}$.

Let S' be a set of 2n points uniformly distributed in the unit square. For each triple (p_i, q_i, r_i) in S' let X_{p_i, q_i, r_i} be the indicator variable having value 1 if the area of the triangle determined by (p_i, q_i, r_i) is less than $\varepsilon = \frac{1}{100n^2}$.

The probability that the area of some specific triangle is less than $\frac{1}{100n^2}$ is less than

$$16\pi\varepsilon = \frac{16\pi}{100n^2} \le \frac{0.6}{n^2}$$

This is also the expected value of X_{p_i,q_i,r_i} .

If X denotes the number of triangles with area less than $\frac{1}{100n^2}$, then

$$\mathbf{E}[X] = \sum_{p,q,r\in S'} \mathbf{E}[X_{p_i,q_i,r_i}] \le \binom{2n}{3} 0.6n^{-2} \le n.$$

Finally, by throwing away an arbitrary vertex from each of such "small area triangles", we are left with a new set S'' of points the expected size of which (of S''), is not less than *n*, in which no small-area-triangles exist.

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Therefore, there exists a set S'', of size *n*, such that $T(S'') \ge \frac{1}{100n^2}$.

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$$\mathbf{E}[Y] = \mathbf{E}[\sum_{e \in E} Y_e] = \sum_{e \in E} \mathbf{E}[Y_e] = \frac{nk}{2}p^2.$$

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After deleting all edges from G_S , by dropping a vertex from each of such edges, it remains a set S^* of the expected size $\mathbf{E}[|S^*|] = \mathbf{E}[|S| - Y]$, and therefore

$$\mathbf{E}[|S|-Y] = \mathbf{E}[|S|] - \mathbf{E}[Y] = np - \frac{nk}{2}p^2.$$

Create a set $S \subset V$ by putting into S each vertex independently with probability p (to be specified later). It therefore holds for the average size of S that: $\mathbf{E}[|S|] = np$. Let G_S be the subgraph of G = (V, E), induced by S.

For any $e \in E$ let Y_e be the indicator variable that has value 1 if $e \in E(G_S)$ - the set of edges of G_S - and 0 otherwise.

 $\mathbf{E}[Y_e] = p^2$ because an edge belongs to $E(G_S)$ iff both if its endpoints are in S, what happens with probability p^2 . Let $Y = |E(G_S)|$. It holds

$$\mathbf{E}[Y] = \mathbf{E}[\sum_{e \in E} Y_e] = \sum_{e \in E} \mathbf{E}[Y_e] = \frac{nk}{2}p^2.$$

After deleting all edges from G_S , by dropping a vertex from each of such edges, it remains a set S^* of the expected size $\mathbf{E}[|S^*|] = \mathbf{E}[|S| - Y]$, and therefore

$$\mathbf{E}[|S| - Y] = \mathbf{E}[|S|] - \mathbf{E}[Y] = np - \frac{nk}{2}p^2$$

The last expression has the largest value for $p = \frac{1}{k}$ and in such a case

$$\mathbf{E}[|S^*|] = \mathbf{E}[|S| - Y] = \frac{n}{2k}.$$

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Finding an explicit construction of OR-concentrators is a non-trivial task. However, the probabilistic method can be used to show the existence of such concentrators.

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We first derive an upper bound on $Pr[\xi_s]$, for any particular fixed *s*, and then we show the upper bound on the sum of $Pr[\xi_s]$ over all $s \leq \alpha n$.

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Theorem: There is an integer n_0 such that for all $n > n_0$ there is an $(n, 18, \frac{1}{3}, 2)$ OR-concentrator.

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Fix any subset $S \subseteq L$ of size *s*, and any subset $T \subseteq R$ of size *cs*.

(There are
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 ways of choosing *S*, and $\begin{pmatrix} n \\ cs \end{pmatrix}$ ways of choosing *T*.)

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Fix any subset $S \subseteq L$ of size *s*, and any subset $T \subseteq R$ of size *cs*.

(There are
$$\binom{n}{s}$$
 ways of choosing *S*, and $\binom{n}{cs}$ ways of choosing *T*.)
The probability that *T* contains all of at most *ds* neighbours of the vertices in *S* is $\left(\frac{cs}{n}\right)^{ds}$.

$$\Pr[\xi_s] \leq \binom{n}{s} \binom{n}{cs} \left(\frac{cs}{n}\right)^{ds}$$

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$$= {\left[\left(\frac{s}{n}\right)^{d-c-1} e^{1+c} c^{d-c}\right]^s}.$$

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Since:

$$r=\left(\frac{2}{3}\right)^{18}\left(3e\right)^3\leq\frac{1}{2}$$

we have

$$\sum_{s>1} \Pr[\xi_s] \le \sum_{s>1} r^s = \frac{r}{1-r} < 1.$$

The probability that there exists an $(n, 18, \frac{1}{3}, 2)$ concentrator is therefore positive.

RANDOMIZED PERMUTATION ROUTING on HYPERCUBES

The first result concerning permutation routing from the previous chapter said that some oblivious routings, for example the so-called left to right routing, are very simple, but they may take expenential time for the delivery of some permutations.

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The second result said that randomized oblivious routing algorithms cn be much more efficient concerning the number of steps. Namely, that there is a randomized oblivious routing algorithm that can route any permutation in 15d steps with probability $1 - \frac{1}{n}$.

As another example, we will show that probabilistic method can be used to prove the existence of a routing algorithm that has as good performance, as the previous one, concerning the number of routing steps and uses much less randomness to do that.

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■ A proof of the existence (by the probabilistic method) of a randomized routing algorithm that uses (within a constant factor) only 3*d* random bits to route *d*-dimensional hypercubes.

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■ A proof of the existence (by the probabilistic method) of a randomized routing algorithm that uses (within a constant factor) only 3*d* random bits to route *d*-dimensional hypercubes.

As a consequence we get that our randomized oblivious routing algorithm, from previous chapter, that used $d2^d$ random bits to route a *d*-dimensional hypercube, uses much too much random bits.

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Note: Every randomized oblivious algorithm can be expressed by sequences

$$(A_1,\ldots,A_r), (p_1,\ldots,p_r),$$

where each A_j is a deterministic oblivious routing algorithm and each p_j is the probability that we use A_j on a run of the randomized routing algorithm.

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Are so many random bits indeed necessary for efficient randomized routing?

Theorem: For every *d* there exists a randomized oblivious scheme (algorithm) for a permutation routing on the hypercube with $n = 2^d$ nodes that uses only 3d random bits and still runs in the expected time 15d at most.

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Proof: Notation We say that a set $\mathcal{B} = \{B_1, \ldots, B_t\}$ of deterministic oblivious permutation routing algorithms on H_d is an efficient routing scheme, if for any input instance, the expected number of steps using a randomly chosen algorithm from \mathcal{B} is at most 15d.

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Our resulting randomized routing scheme will randomly choose n^3 of n^n possible deterministic oblivious routing algorithms. (n^n is due to the fact that there are *n* sources and for each one we can choose from *n* possible intermediate destinations.)

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Let us denote such deterministic algorithms by A_j , $1 \le j \le n^n$. On an *n*-node network there are *n*! distinct possible instances of the permutation routing problem, one for each permutation on $\{1, 2, ..., n\}$.

By our randomized routing result: (with probability at least $1 - \frac{1}{n}$ every packet reaches its destination in 14*n* or fewer steps) for any particular π_i a fraction of at most $\frac{1}{n}$ of the algorithms A_j are bad - which of them are bad may differ from instance to instance.

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Experiment: Choose n^3 indices i_1, \ldots, i_{n^3} , randomly, independently and uniformly from the set $\{1, \ldots, n^n\}$. We show that the set of deterministic algorithms

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is an efficient routing scheme with a positive probability. This will imply that an efficient routing scheme exists for any $n = 2^d$.

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For any π_i , a fraction of at most $\frac{1}{n}$ of the algorithms $\mathcal{A}_1, \ldots, \mathcal{A}_{n^n}$ is bad. Therefore, the expected number of algorithms in \mathcal{A} that are bad for π_i is at most $n^3 \cdot \frac{1}{n} = n^2$. Let the indicator variable X_j be 1 if A_{i_j} is bad, $1 \le j \le n^3$, and 0 otherwise.

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$$\Pr[X \ge (1+\delta)\mu] \le F^+(\mu,\delta) < e^{-rac{\mu\delta^2}{4}}$$

(Chapter 6, page 10), to obtain (for $\mu = n^2, \delta = 1$) that the probability that more than $2n^2$ of the algorithms in \mathcal{A} are bad for π_i is $\leq e^{\frac{-n^2}{4}}$. Let \mathcal{B}_i denote the **bad** event that more than $2n^2$ algorithms in \mathcal{A} are bad for π_i .

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Let us denote this subset $\mathcal{B} = \{B_{i_1}, \ldots, B_{i_{n^3}}\}$. \mathcal{B} is an efficient routing scheme: for any π_i a randomly chosen algorithm form \mathcal{B} fails to route π_i within 14*n* steps with probability at most $\frac{2n^2}{n^3} = \frac{2}{n}$.

THE LOVÁSZ LOCAL LEMMA

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Lovász Local lemma handle the situation for the case where events are generally not independent of each other, but each collection of events that are not independent of some particular event A has low total probability.

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Later, it has been shown that when events are determined by some underlying sets of independent variables and independence between two events is detected by having non-overlapping sets of underlying variables, an actual solution can be found in polynomial expected time

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Then $Pr\left(\bigcap_{i=1}^{n} \overline{E}_{i}\right) > 0.$

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Theorem: If, for a fixed k, any path in any F_i shares edges with no more than k paths in any F_j , $j \neq i$, and $8nk/m \leq 1$, then there is a way to choose n edge-disjoint paths connecting the n given pairs

PROOF of THEOREM

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Then

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PAUL ERDÖS

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STORIES about **ERDÖS**

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Paul Erdös in a letter to Vera Sós

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 - added "let us now look for a Book proof".
- In 1985 he started his lecture in a math camp by saying: "You do not have to believe in the God, but you should believe in The Book".